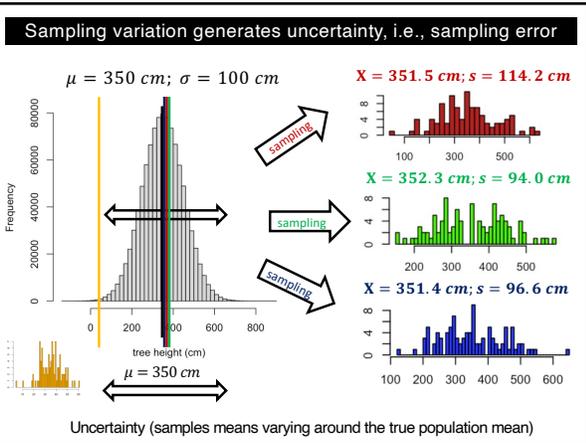


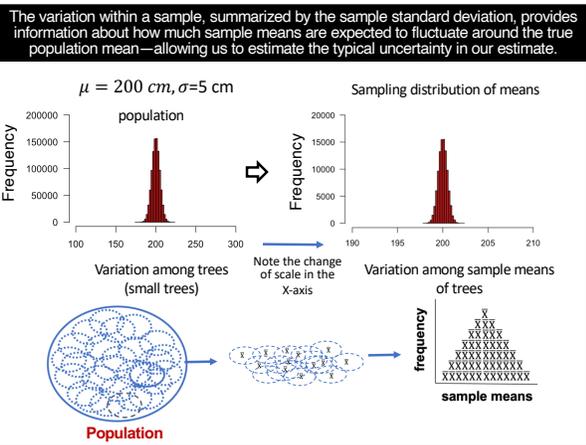
Lecture 10: Estimating with uncertainty, but with a degree of certainty (i.e., with some confidence), part 2

Key statistical concepts for understanding confidence intervals, statistical procedures, and statistical reasoning

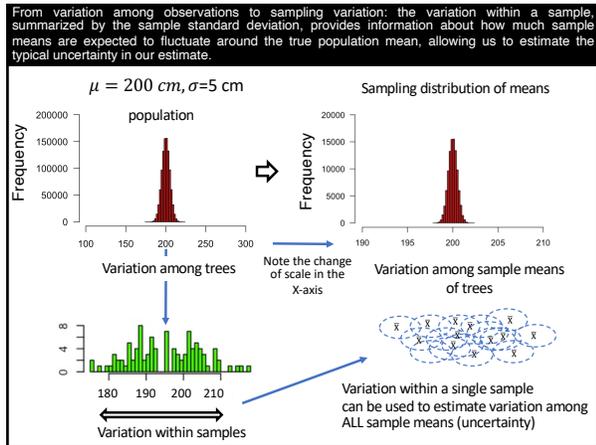
1



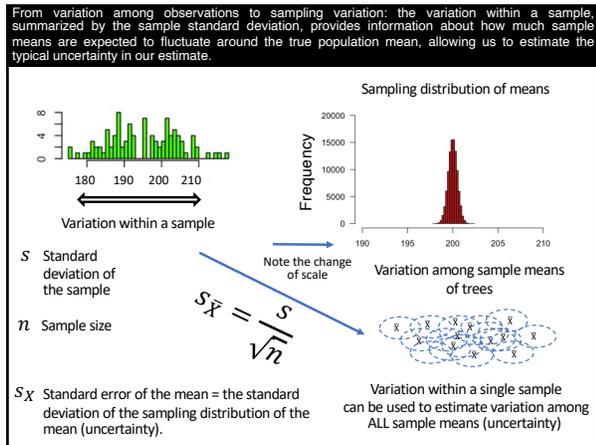
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3



4



5

$S_x = \frac{s}{\sqrt{n}}$

THE STANDARD ERROR OF THE MEAN

Standard deviation (s): If you measure 25 students and the SD is 10 cm, it means a typical student's height differs from the sample average by about 10 cm.

If you repeatedly (randomly) sampled different groups of 25 students, their averages would vary slightly around the true population mean. And the same for any number of students that build a sample.

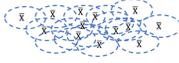
$S_x = \frac{10}{\sqrt{25}} = 2 \text{ cm}$

The 2 cm is the standard error and is a measure of uncertainty. It tells you: if you repeatedly took random samples of 25 students and calculated their average height, those averages would typically differ from the true population mean by about 2 cm.

UNCERTAINTY: s_x (the standard error) estimates how wrong our sample mean is expected to be, on average, from the true population mean—purely because we relied on a sample.

6

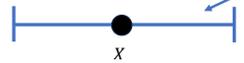
CONFIDENCE INTERVAL FOR THE MEAN

$$s_X = \frac{10}{\sqrt{25}} = 2\text{cm}$$


UNCERTAINTY: s_X (the standard error) estimates how wrong our sample mean is expected to be, on average, from the true population mean—purely because we relied on a sample.

Confidence interval = $\bar{X} \pm$
quantity * s_X

Very plausible (high confidence) that the population parameter μ is somewhere within the 95% confidence interval.



The quantity (i.e., which establishes the critical value or margin of error) varies with the confidence level we choose (e.g., 95% or 99%).

7

Although we have not yet shown how to estimate a confidence interval, we already know that it depends on the standard error of the mean, a measure of how much sample means vary due to sampling.

The standard deviation of the sampling distribution σ_Y is called standard error (SE) and is exactly:

$$\sigma_Y = \sqrt{\lim_{N \rightarrow \infty} \sum_{i=1}^N \frac{(X_i - \mu)^2}{N}} = \frac{\sigma}{\sqrt{n}}$$

N is all infinite samples and n is sample size

The left side is a definition using infinite repetition and the right side is a mathematical consequence of the probability model. Statistics closes these gaps.

$\sqrt{\lim_{N \rightarrow \infty} \sum_{i=1}^N \frac{(X_i - \mu)^2}{N}}$ → Not observable: infinitely many samples

$\frac{\sigma}{\sqrt{n}}$ → closed-form result from probability theory. variability of sample means depends only on: the population standard deviation σ , and the sample size n .

[REVISED]

8

Although we have not yet shown how to estimate a confidence interval, we already know that it depends on the standard error of the mean, a measure of how much sample means vary due to sampling.

Given that we almost never know the population standard deviation, we estimate it with the sample value based on the sample standard error:

$$\sigma_X = \frac{\sigma}{\sqrt{n}} \quad \leftarrow \quad s_X = \frac{s}{\sqrt{n}}$$

σ_X = the standard deviation of the sampling distribution of means (standard error); σ = the standard deviation of the population.

The standard error of the mean, SE_X , estimates the standard deviation of the sampling distribution of the mean; that is, how much sample means are expected to vary around the true population mean across repeated samples. Remarkably, this quantity can be estimated from a single sample (a principle we will examine in detail in Tutorial 5).

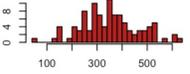
[REVISED]

9

How to calculate a confidence interval in practice:

$$S_X = \frac{s}{\sqrt{n}} = \frac{114.2}{\sqrt{100}} = 11.42 \text{ cm}$$

$X = 351.5 \text{ cm}; s = 114.2 \text{ cm}$



S_X estimates the standard deviation of the sampling distribution of the mean (uncertainty).

Based on this sample, we estimate that sample means typically differ from the true population value by about 11.42 cm due to sampling variability.

Confidence interval = $351.5 \text{ cm} \pm$
quantity * 11.42 cm

Obviously, a different sample will generate a different estimate of this error.

[REVISED]

10

Sampling distribution of the mean, variance, standard deviation, and standard error obtained from the population (1, 2, 3, 4, 5), considering all possible samples drawn with replacement for sample sizes $n = 2, 3, 4$.

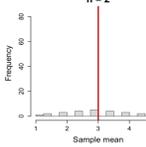
Obs 1	Obs 2	Sample mean	Mean - 3	(Mean - 3) ²	Sample Var (n-1)	Sample SD (n-1)	Sample SD / $\sqrt{2}$
1	1	1.0	-2.0	4.00	0.000	0.000	0.000
1	2	1.5	-1.5	2.25	0.500	0.707	0.500
1	3	2.0	-1.0	1.00	2.000	1.414	1.000
1	4	2.5	-0.5	0.25	4.500	2.121	1.500
1	5	3.0	0.0	0.00	8.000	2.828	2.000
2	1	1.5	-1.5	2.25	0.500	0.707	0.500
2	2	2.0	-1.0	1.00	0.000	0.000	0.000
2	3	2.5	-0.5	0.25	0.500	0.707	0.500
2	4	3.0	0.0	0.00	2.000	1.414	1.000
2	5	3.5	0.5	0.25	4.500	2.121	1.500
3	1	2.0	-1.0	1.00	2.000	1.414	1.000
3	2	2.5	-0.5	0.25	0.500	0.707	0.500
3	3	3.0	0.0	0.00	0.000	0.000	0.000
3	4	3.5	0.5	0.25	0.500	0.707	0.500
3	5	4.0	1.0	1.00	2.000	1.414	1.000
4	1	2.5	-0.5	0.25	4.500	2.121	1.500
4	2	3.0	0.0	0.00	2.000	1.414	1.000
4	3	3.5	0.5	0.25	0.500	0.707	0.500
4	4	4.0	1.0	1.00	0.000	0.000	0.000
4	5	4.5	1.5	2.25	0.500	0.707	0.500
5	1	3.0	0.0	0.00	8.000	2.828	2.000
5	2	3.5	0.5	0.25	4.500	2.121	1.500
5	3	4.0	1.0	1.00	2.000	1.414	1.000
5	4	4.5	1.5	2.25	0.500	0.707	0.500
5	5	5.0	2.0	4.00	0.000	0.000	0.000

11

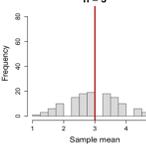
Sampling distribution of the mean obtained from the population (1, 2, 3, 4, 5), considering all possible samples drawn with replacement for sample sizes $n = 2, 3, 4$.

True population mean = mean of sample means

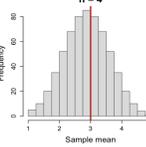
n = 2



n = 3



n = 4



One critical observation is that **the mean of all possible sample means is exactly the population mean**. This is important because it tells us that, on average, the sampling process is **unbiased**: repeated sampling does not systematically overestimate or underestimate the true population value.

Even though individual samples can yield means that are far from the population mean, especially when sample size is small, these deviations balance out across all possible samples.

12

Recall from lecture 6: The mean can be understood as the center of gravity of a distribution: the point at which the values on either side balance each other.

Sum (left) = -7.5 Sum (right) = 7.5

$\bar{X} = 4.5$

Sum (left) + Sum (right) = -7.5 + 7.5 = 0

13

Recall from lecture 5: the sum of deviations from the mean is always zero, making the mean the 'center of gravity' of a distribution, i.e., the values on either side of the mean balance each other.

Observations (Y_i)	Deviations ($Y_i - \bar{Y}$)
0.9	-0.475
1.2	-0.175
1.2	-0.175
1.3	-0.075
1.4	0.025
1.4	0.025
1.6	0.225
2.0	0.625
Sum	0.000

$\bar{X} = 4.5$

Whitlock & Schluter, The Analysis of Biological Data, 3e © 2020 W. H. Freeman and Company

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Sampling distribution: the **sample means** on either side of the **true population mean μ** balance each other, i.e., the sum = 0.

sampling error

Obs 1	Obs 2	Sample means	Sample means - μ
1	1	1.0	-.2
1	2	1.5	-.1
1	3	2.0	0
1	4	2.5	.1
1	5	3.0	.2
2	1	1.5	-.1
2	2	2.0	0
2	3	2.5	.1
2	4	3.0	.2
2	5	3.5	.3
3	1	2.0	0
3	2	2.5	.1
3	3	3.0	.2
3	4	3.5	.3
3	5	4.0	.4
4	1	2.5	.1
4	2	3.0	.2
4	3	3.5	.3
4	4	4.0	.4
4	5	4.5	.5
5	1	3.0	.2
5	2	3.5	.3
5	3	4.0	.4
5	4	4.5	.5
5	5	5.0	.6
MEAN		3.0	0

The mean of all possible sample means is exactly the population mean.

The sampling process of the mean is unbiased: repeated sampling does not systematically overestimate or underestimate the true population value, i.e., they balance each other out.

n = 2

15

The expected value (E) is the average of the sampling distribution; that is, the mean value we would obtain if we could repeat the sampling process indefinitely.

$E(X) = \mu$ The expected value of the sample mean equals the population mean (unbiased).

$E(s^2) = \sigma^2$ The expected value of the sample variance equals the population variance (biased).

$E(s) \neq \sigma$ The expected value of the sample standard deviation IS NOT equal the population standard deviation (biased).

$E\left(\frac{s}{\sqrt{n}}\right) \neq \frac{\sigma}{\sqrt{n}}$ The expected value of the sample standard error IS NOT equal the population standard error (biased).

E is called the expectation operator. It represents the theoretical long-run (infinite) average of a random variable.

22

Although we have not yet shown how to estimate a confidence interval, we already know that it depends on the standard error of the mean, a measure of how much sample means vary due to sampling.

Given that we almost never know the population standard deviation, we estimate it with the sample value based on the sample standard error:

$\sigma_X = \frac{\sigma}{\sqrt{n}}$ $s_X = \frac{s}{\sqrt{n}}$

σ_X = the standard deviation of the sampling distribution of means (standard error) ; σ = the standard deviation of the population.

?

*Confidence interval = $X \pm$ quantity * s_X*

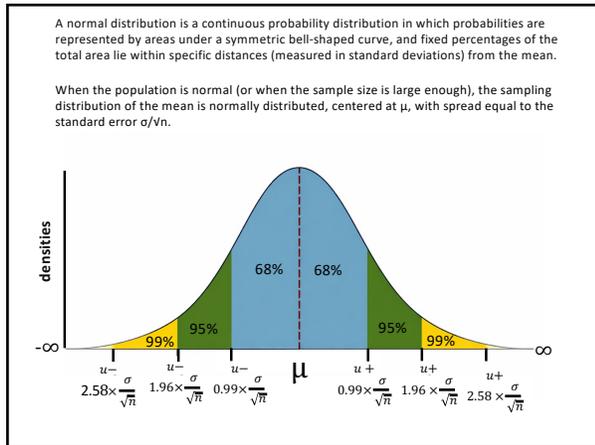
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23

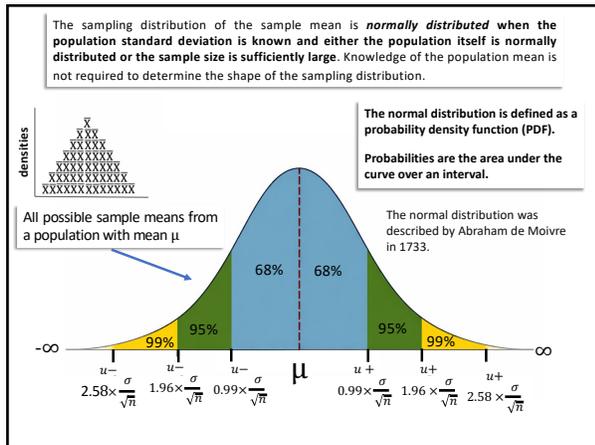
The logic behind confidence intervals before we see how the bias in the sample standard deviation is corrected and why that correction matters.



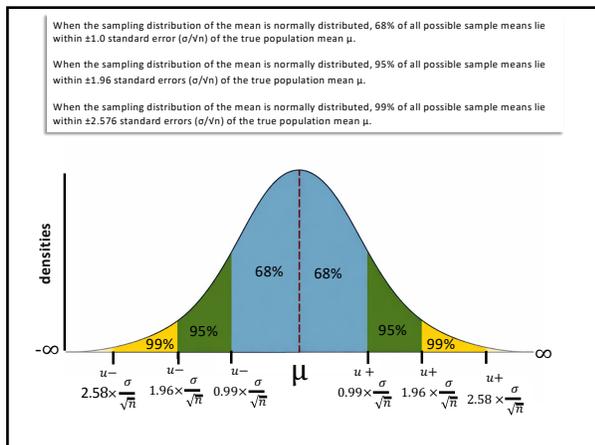
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25



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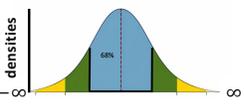


27

The sampling distribution of the means goes from $-\infty$ to ∞

How many possible samples of 100 trees out of 100000 trees?
1e+15 (zeros)

How many possible samples of 100 trees out of 1000000?
10768272362e+432 (zeros)



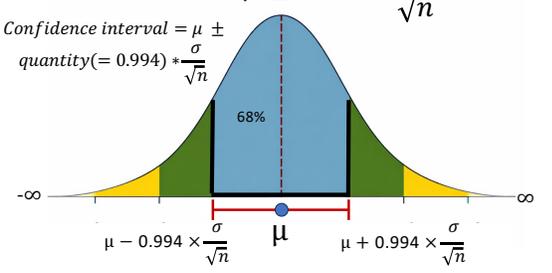
The **human body** consists of some 37.2 trillion **cells** (3.72e+13 zeros)

28

68% of all possible sample means (i.e., 68% of the area under the sampling distribution of the mean) lies within $\mu \pm 0.994 \times \sigma$, i.e., there is a 68% probability that a randomly drawn sampling from a population with mean μ and standard deviation σ will be within $0.994 \times \frac{\sigma}{\sqrt{n}}$ of the true mean, i.e., within the interval.

68% CI: $\mu \pm 0.994 \times \frac{\sigma}{\sqrt{n}}$

Confidence interval = $\mu \pm$ quantity(= 0.994) $\times \frac{\sigma}{\sqrt{n}}$

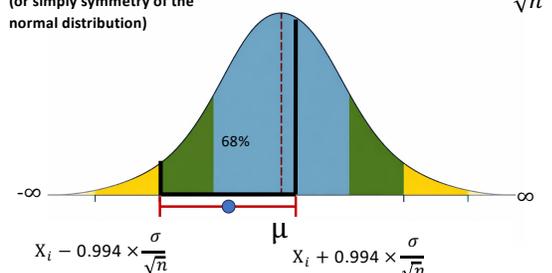


29

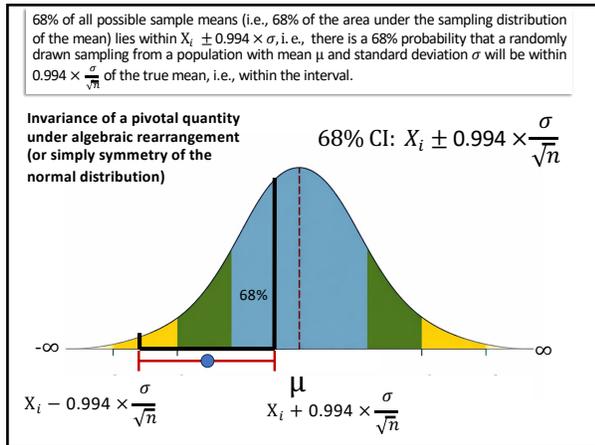
68% of all possible sample means (i.e., 68% of the area under the sampling distribution of the mean) lies within $X_i \pm 0.994 \times \sigma$, i.e., there is a 68% probability that a randomly drawn sampling from a population with mean μ and standard deviation σ will be within $0.994 \times \frac{\sigma}{\sqrt{n}}$ of the true mean, i.e., within the interval.

Invariance of a pivotal quantity under algebraic rearrangement (or simply symmetry of the normal distribution)

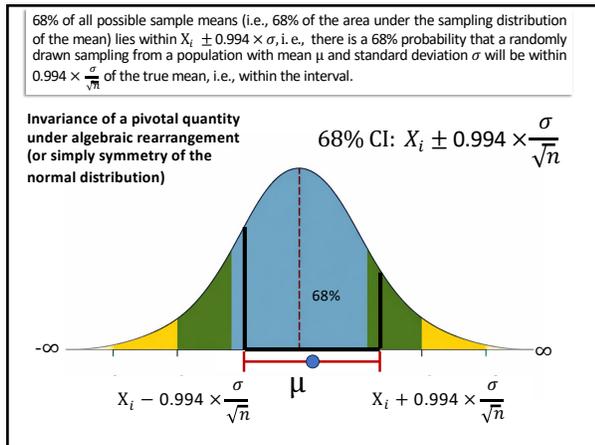
68% CI: $X_i \pm 0.994 \times \frac{\sigma}{\sqrt{n}}$



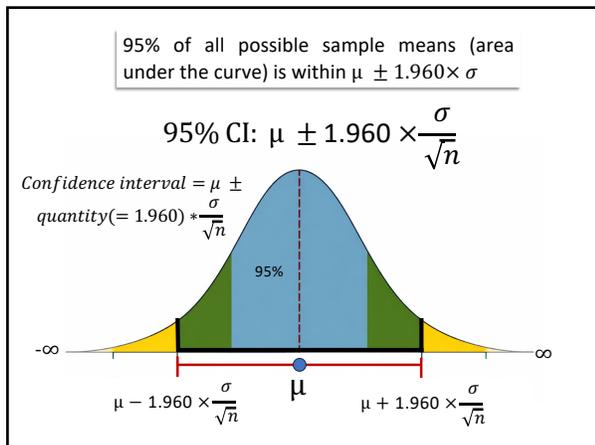
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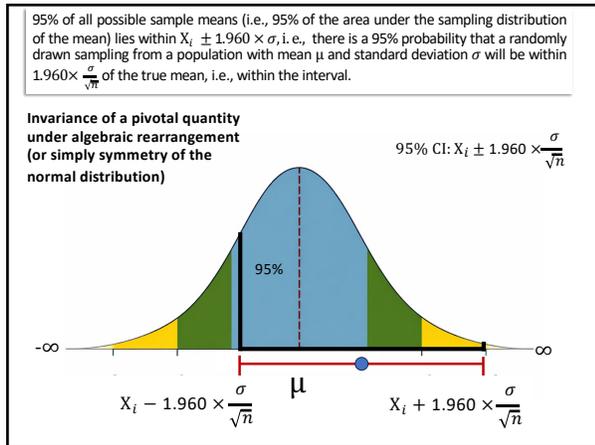
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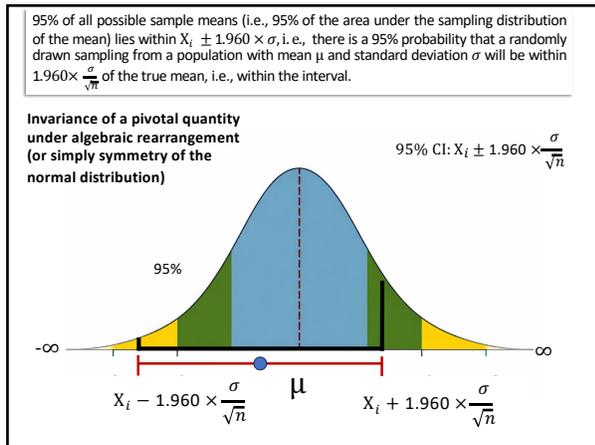
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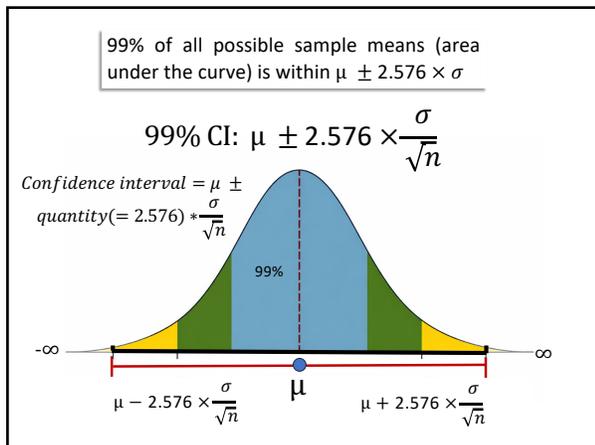
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34



35



36

Remember that the sample standard deviation that, on average, tends to underestimate the true population standard deviation. It underestimates even more when sample sizes are smaller (Jensen's inequality).

So, what happens if we replace the population standard deviation with the sample standard deviation when building our interval still assuming a normal distribution for the sampling distribution of the mean?

$CI = \mu \pm 1.960 \cdot \frac{S_i}{\sqrt{n}}$ This interval won't work because the sample-based standard deviation is biased.

For a 95% confidence interval, we expect that only 5% of intervals will fail to include the true population mean.

However, if the sample standard deviation underestimates the true population standard deviation, especially when the sample size is small (Jensen's inequality), the resulting intervals will, on average, be too narrow.

As a consequence, more than 5% of such intervals will fail to contain the true population mean.

[REVISED]

40

However, if the sample standard deviation underestimates the true population standard deviation, especially when the sample size is small (Jensen's inequality), the resulting intervals will, on average, be too narrow.

As a consequence, more than 5% of such intervals will fail to contain the true population mean.

— True population standard deviation (1.414214)
— Mean of sample standard deviations (n=2: 1.131371; n=3: 1.287381; n=4: 1.340293)

$E[\text{width}_{\text{sample-based}}] < E[\text{width}_{\text{population-based}}]$

Sample-based (unknown σ) $CI = X_i \pm 1.960 \cdot \frac{S_i}{\sqrt{n}}$ \rightarrow $CI = X_i \pm 1.960 \cdot \frac{\sigma}{\sqrt{n}}$ Population-based (known σ)

[REVISED]

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How to adjust confidence intervals for the biases in sample-based standard deviations?

42

Now imagine the following computational approach (which is approximated with calculus in practice):

- 1) We repeatedly (10,000,000) take samples of size n from a normally distributed population
- 2) For each sample, we calculate the sample mean \bar{X}_i the sample standard error of the mean $s_{\bar{X}_i} = \frac{s_i}{\sqrt{n}}$.
- 3) We then subtract the sample \bar{X}_i by the population mean μ and divide the result by the sample standard error of the mean.
- 4) Let's call this quantity the t value for the i^{th} sample: $t_i = \frac{\bar{X}_i - \mu}{\frac{s_i}{\sqrt{n}}}$
- 5) Combine (pool) all 10,000,000 computed t-values, one from each generated sample, and plot their distribution as a probability density function (e.g., a histogram scaled to density or a smooth density curve). The resulting density is the t-distribution (with degrees of freedom (later in the course) = $n-1$, because s_i is estimated as:

$$s = \sqrt{\frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n-1}}$$

43

The t-distribution - developed analytically by *student* (William Sealy Gosset) in 1908.

44

Some t-values are negative (when sample mean is smaller than the population mean μ) and others are positive (when sample mean is greater than μ).

What is the mean of the t-distribution?

Obs 1	Obs 2	Sample means	Sample means - μ
1	1	1.0	-2
1	2	1.5	-1.5
1	3	2.0	-1
1	4	2.5	-0.5
1	5	3.0	0
2	1	1.5	-1.5
2	2	2.0	-1
2	3	2.5	-0.5
2	4	3.0	0
2	5	3.5	0.5
3	1	2.0	-1
3	2	2.5	-0.5
3	3	3.0	0
3	4	3.5	0.5
3	5	4.0	1
4	1	2.5	-0.5
4	2	3.0	0
4	3	3.5	0.5
4	4	4.0	1
4	5	4.5	1.5
5	1	3.0	0
5	2	3.5	0.5
5	3	4.0	1
5	4	4.5	1.5
5	5	5.0	2
MEAN		3.0	0

For any given sample i , its confidence interval is:

$$CI_i = \bar{X}_i \pm t * \frac{s_i}{\sqrt{n}}$$

[REVISED]

45

Because we subtract the true population mean, the t-statistic measures the deviation of the sample mean from μ , regardless of the numerical value of μ .

→ This removes location: the distribution is centered at zero.

Because we divide by the sample standard error $\frac{S_i}{\sqrt{n}}$, the deviation is expressed in units of estimated variability.

→ This removes scale: the statistic becomes unit-free and comparable across populations.

What remains is a standardized quantity whose distribution depends only on sample size (or more precisely, degrees of freedom), not on μ or σ .

The same t value (e.g., ± 2.365 for $df=7$) can be used regardless of whether $\mu = 10, 100, \text{ or } 10,000$ or what the standard deviation is.

→ The t-distribution is UNIVERSAL and only varies as a function of sample size!

46

The sampling distribution of means that varies as a function of the sample size (here $df = \text{degrees of freedom; } df = n - 1$) is t-distributed.

The exact multiplier applied to the sample standard error to construct a 95% confidence interval based on the t-distribution changes with sample size (df).

Sampling distribution of the mean
(normally distributed – infinite sample sizes)

All intervals with the same width – single σ

XX% CI: $X_i \pm Z \times \frac{\sigma}{\sqrt{n}}$

95% CI ($Z = 1.960$): $X_i \pm 1.960 \times \frac{\sigma}{\sqrt{n}}$

Sampling distribution of the mean
(t-distributed – non-infinite sample sizes)

Intervals differ in width – sample-based

XX% CI: $X_i \pm t.df \times \frac{S_i}{\sqrt{n}}$

95% CI: $X_i \pm t.df \times \frac{S_i}{\sqrt{n}}$

[REVISED]

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By now, we have seen in this lecture three types of confidence intervals:

$CI(95\%) = X_i \pm 1.960 \times \frac{\sigma}{\sqrt{n}}$

$CI(95\%) = X_i \pm 1.960 \times \frac{S_i}{\sqrt{n}}$

$E[\text{width}_{\text{sample based}}] < E[\text{width}_{\text{population based}}]$

$CI(95\%) = X_i \pm t.df \times \frac{S_i}{\sqrt{n}}$

95% Confidence interval (CI) for the population mean based on any sample mean \bar{X}_i with known population standard deviation σ . This is correct but we never really know σ in applied situations.

95% Confidence interval (CI) for the population mean based on any sample mean \bar{X}_i with unknown population standard deviation σ and, therefore, based on its (sample) standard deviation S_i . This is incorrect leading to small coverage than anticipated (i.e., less than 95% of the CIs will cover the true population mean. This was shown for pedagogical purposes only, i.e., to show that we need another solution because this won't work. v

95% Confidence interval (CI) for the population mean based on any sample mean \bar{X}_i with unknown population standard deviation σ and, therefore, based on its (sample) standard deviation S_i . This is correct because uses the t-distribution and, as such, won't be biased when using the sample standard deviation S_i .

[NEW]

48

The exact multiplier applied to the sample standard error to construct a 95% confidence interval depends on the sample size.

The sampling distribution of means that varies as a function of the sample size (here $v = \text{degrees of freedom}; v = n - 1$) is t-distributed.

$v = \infty \rightarrow t = Z \rightarrow$ the t distribution becomes normally distributed when sample size is infinite.

95% CI: $X_i \pm t_{df} \times \frac{\sigma}{\sqrt{n}}$

Extreme t-values occur more often than they would under the normal model. Graphically, the curve looks slightly flatter in the center and thicker in the tails. That is what "heavier tails" means: a higher probability of observing values far from zero compared to the normal distribution.

[NEW]

49

$v = \infty \rightarrow t = Z \rightarrow$ the t distribution becomes normally distributed when sample size is infinite.

95% CI: $X_i \pm t_{df} \times \frac{\sigma}{\sqrt{n}}$

95% CI ($Z = 1.960$): $X_i \pm 1.960 \times \frac{\sigma}{\sqrt{n}}$

df	95%	t	Difference from 1.96
5		2.571	+0.611
10		2.228	+0.268
20		2.086	+0.126
30		2.042	+0.082
50		2.009	+0.049
100		1.984	+0.024
∞		1.960	0

Interval width

95% CI: $X_i \pm 2.571 \times \frac{s}{\sqrt{n}}$

95% CI: $X_i \pm 2.086 \times \frac{s}{\sqrt{n}}$

95% CI: $X_i \pm 2.009 \times \frac{s}{\sqrt{n}}$

95% CI: $X_i \pm 1.960 \times \frac{s}{\sqrt{n}}$

As the sample size becomes very large, the t-distribution converges to the normal distribution.

[REVISED]

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Using the wrong distribution affects coverage: If we use the normal distribution instead of t (when σ is unknown), intervals are too narrow and more than 5% of 95% CIs will miss the true mean (undercoverage). With the correct approach: t-based CIs account for extra uncertainty and 5% will miss the true mean in the long run, as intended.

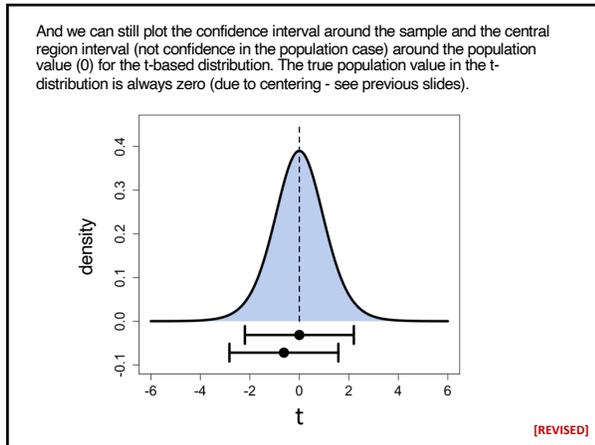
$E[\text{width}_{\text{Normal based}}] < E[\text{width}_{\text{t based}}]$

Normal based: $v\text{CI} = \mu \pm 1.960 \times \frac{s_i}{\sqrt{n}}$

t based: $\text{CI} = \mu \pm t \times \frac{s_i}{\sqrt{n}}$

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How to find the appropriate values of t?
 the old days of tables allow to understand the principle – in practice (today) we use software (e.g., R).

Degrees of freedom (n-1)	Two-sided	50%	60%	70%	80%	90%	95%	98%	99%	99.5%	99.8%	99.9%
1	1.000	1.376	1.963	3.078	6.314	12.71	31.82	63.66	127.3	318.3	636.6	
2	0.816	1.080	1.386	1.886	2.920	4.303	6.965	9.925	14.09	22.32	31.60	
3	0.765	0.978	1.250	1.638	2.353	3.182	4.541	5.841	7.453	10.21	12.92	
4	0.741	0.941	1.190	1.533	2.132	2.776	3.747	4.604	5.598	7.173	8.610	
5	0.727	0.920	1.156	1.476	2.015	2.571	3.365	4.032	4.773	5.893	6.869	
6	0.718	0.906	1.134	1.440	1.943	2.447	3.143	3.707	4.317	5.208	5.959	
7	0.711	0.896	1.119	1.415	1.895	2.365	2.998	3.499	4.029	4.785	5.408	
8	0.706	0.889	1.108	1.397	1.880	2.306	2.896	3.355	3.833	4.501	5.041	
9	0.703	0.883	1.100	1.383	1.833	2.262	2.821	3.250	3.690	4.297	4.781	
10	0.700	0.879	1.093	1.372	1.812	2.228	2.764	3.169	3.581	4.144	4.587	
11	0.697	0.876	1.088	1.363	1.796	2.201	2.716	3.106	3.497	4.025	4.437	
12	0.695	0.873	1.083	1.356	1.782	2.179	2.681	3.055	3.428	3.930	4.318	

Assume a sample size of $n = 9$, then the degrees of freedom would be 8 for the t value to calculate the confidence interval for the sample mean.

$$s = \sqrt{\frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n - 1}}$$

$$\bar{X}_i \pm t_{df} \frac{s_i}{\sqrt{n}} \therefore X_i \pm 2.306 \frac{s_i}{\sqrt{9}}$$

Degrees of freedom are commonly denoted as df (or sometimes ν , the Greek letter nu).

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Let's consider a biological example: The stalk-eyed fly – the span in millimeters of nine male individuals are as follows:

8.69 8.15 9.25 9.45 8.96 8.65 8.43 8.79 8.63 n=9 flies

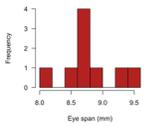
Let's estimate the **95%** confidence interval for the population mean of eyes spans:

$\bar{X} = 8.778$ $s = 0.398$

$SE_{\bar{X}} = \frac{0.398}{\sqrt{9}} = 0.133$

$t_{0.05(2),8} = 2.306$

$\bar{X} - 2.306 \times 0.133 < \mu < \bar{X} + 2.306 \times 0.133$
 8.47 mm < μ < 9.08 mm




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In practice (modern days), we use software

$\bar{X} = 8.778$ $s = 0.398$ Let's estimate the **95%** confidence interval for the population mean of eyes spans:

$SE_{\bar{X}} = \frac{0.398}{\sqrt{9}} = 0.133$

$t_{0.05(2),8} = 2.306$

$\bar{X} - 2.306 \times 0.133 < \mu < \bar{X} + 2.306 \times 0.133$
 8.47 mm < μ < 9.08 mm



```
> t.test(stalkie$eyespan, conf.level = 0.95)$conf.int
[1] 8.471616 9.083940
attr(,"conf.level")
[1] 0.95
```

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In practice (modern days), we use software

$\bar{Y} = 8.778$ $s = 0.398$ Let's estimate the **99%** confidence interval for the population mean of eyes spans:

$SE_{\bar{Y}} = \frac{0.398}{\sqrt{9}} = 0.133$

$t_{0.05(2),8} = 3.355$

$\bar{Y} - 3.355 \times 0.133 < \mu < \bar{Y} + 3.355 \times 0.133$
 8.33 mm < μ < 9.22 mm

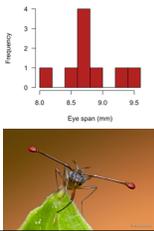


```
> t.test(stalkie$eyespan, conf.level = 0.99)$conf.int
[1] 8.332292 9.223264
attr(,"conf.level")
[1] 0.99
```

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Comparing the two intervals, the implications for precision depend on the biological context and the level of certainty required. In practice, we most often report the 95% confidence interval, which provides a balance between coverage and precision—offering less coverage than a 99% interval but greater precision (narrower width).

95% confidence interval: $8.47 \text{ mm} < \mu < 9.08 \text{ mm}$
 99% confidence interval: $8.33 \text{ mm} < \mu < 9.22 \text{ mm}$



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Confidence intervals can be used in many situations; e.g., decision making

The optimal average body length for a healthy brook trout population is 24 cm or greater. According to management policy, if the true population mean length is below 24 cm, the lake must be closed to fishing.

At the beginning of the fishing season, a government biologist surveys a heavily exploited lake, where strict regulation is particularly important. The biologist captures, anesthetizes, and measures 200 brook trout. The sample mean length is 21.3 cm, with a sample standard deviation of 3.2 cm.

Should the lake be opened for fishing?

$$Y \pm t \times SE_{Y_s} \therefore 21.2 \pm 1.971957 \times \frac{3.2}{\sqrt{200}} = 21.2 \pm 0.4462$$

20.8cm $\bar{X}=21.2\text{cm}$ 21.6cm


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