

Dealing with "some" important statistical assumptions.

#### 1) The issue of normality (today):

- Parametric (e.g., ANOVA): assume parametrized families of probability distributions (e.g., normal defined by two parameters, i.e., mean and variance). Parameter estimates tend to be sensitive to nonnormality (e.g., issue in regression slopes), but not necessarily in statistical hypothesis testing (P-values may be not as sensitive).

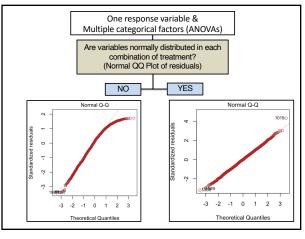
- Non-parametric: either distribution free (e.g., permutation tests) or ranked based tests.

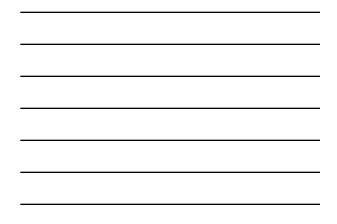
2

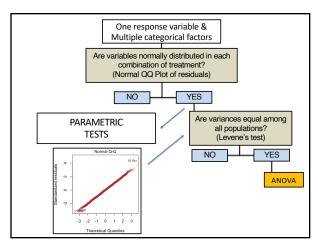
Dealing with "some" important statistical assumptions.

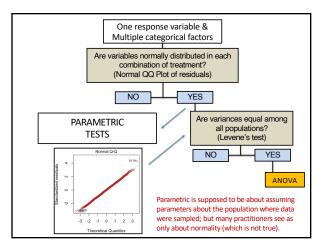
# 2) The issue of homogeneity of variances (later in the course):

- Standard (e.g., ANOVAs, regressions) assume homoscedasticity.
- Robust approaches (Welch's ANOVA, Weighted least squares) are good to deal with heteroscedasticity.

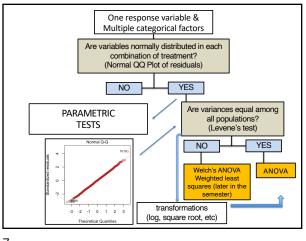




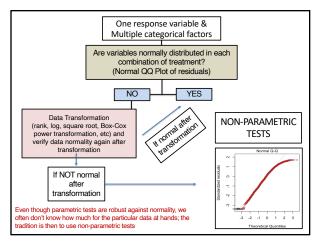




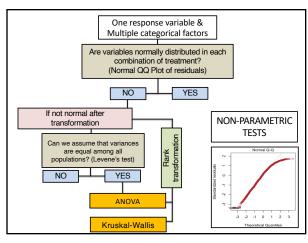




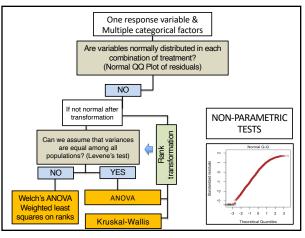


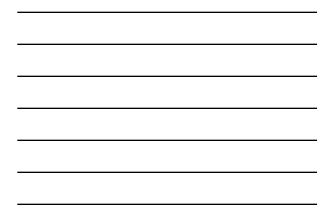


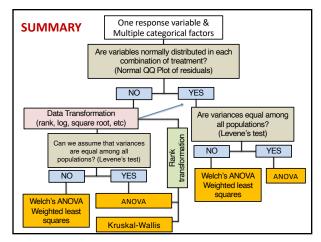




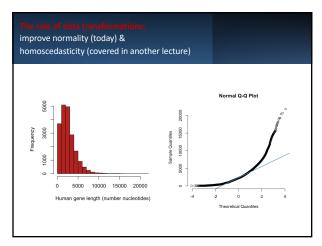




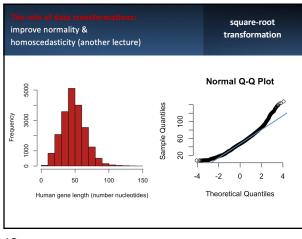




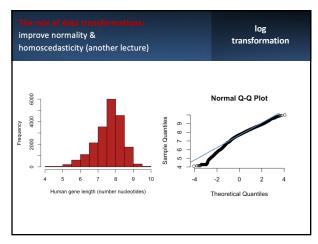














14

#### A few words on data transformation

One size may not fit all:

1) One transformation may help approximate normality, but another transformation may be required to approximate homoscedasticity (e.g., log(sqrt(data))).

2) One transformation may negate (reverse) the other – the one that makes the data approximate homoscedasticity may make data non-longer normal.

3) If data are complex (e.g., several predictors in a regression model), it may not be possible that one single transformation will allow data to behave properly under assumptions. Possible solution: focus on analytical solutions (many covered in this course) and not always transformations; or combine different

15

transformation.



3) If data are complex (e.g., several predictors in a regression model), it may not be possible that one single transformation will allow data to behave properly under assumptions.

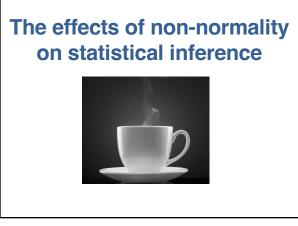
Possible solution: focus on analytical solutions (many covered later in the semester) and not always transformations; or combine different transformation.

The R Package trafo for Transforming Linear Regression Models

Lily Medina Piedad Castro Humbohl Universität zu Berlin Ann-Kristin Kreutzmann Preie Universität Berlin

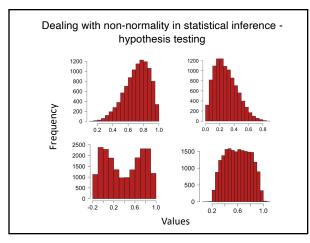
Abstract The timer regrams model also we addy and for densitytics, prediction, and addy approxes. The second relative was addressed on the second second second conduct regrams and table of nonlong and the second second second conduct regrams and table of nonlong answering the second secon

16

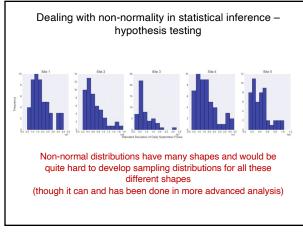




17







#### The effects of non-normality on statistical test

Parametric tests assuming normality (e.g., t-test & ANOVA) are often robust against non-normality; but depending on the type of non-normality (shape), parametric tests can have type I errors different (often greater) from alpha; and also low power (increased type II error).

One challenge is to separate normality from heteroscedasticity issues (even in simulations).

The other challenge is when samples come all from populations with different distributions (even though they could have the same means, i.e.,  $H_0$  is true).

20

#### The effects of non-normality on statistical test

Parametric tests assuming normality (e.g., t-test & ANOVA) are often robust against non-normality; but depending on the type of nonnormality (shape of the distribution), parametric tests can have type I errors (false positives) that differ (often greater) from alpha; and low power (increased type II error; false negatives).

101. 2013 May;66(2):224-44. doi: 10.1111/j.2044-8317.2012.02047.x. Epub 2012 May 24. Br J Math Stat Pa The impact of sample non-normality on ANOVA and alternative methods. Lantz B<sup>1</sup>. Author information

Abstract

Abstract In this journal, Caromerman (2004, 2011) has discussed preliminary tests that researchers often use to choose an appropriate method for comparing locations when the essumption of normality is solutiful. The conceptual problem with this approach is that auch a two-stage process makes both the power and the significance of the entire procedue uncertain, supple and type II errors are possible at both stages. A speal entror at the first stage, for example, will obviously increases the probability of a bype II error at the second stage. Based on the loss of Solumider of at (2010), which proposes that invalided and and annual test data to indeed with interposed to the dispart of normality in the propose Bound Ford at (2010), which proposes that invalided and an advect data to indeed with interposed to the dispart of normality. This paper Bound Ford at (2010), which proposes that invalided and an advect data to indeed with interposed to the dispart of normality. This paper Bound Ford at (2010), which proposes that invalided and an advect data to indeed with entropose taxies. Board or formality in the paper Bound Ford at (2014), which proposes that considerably less sensitive to the degree of apple taxies. The overall Bound Ford at (2014), which proposes that considerably less sensitive to the degree of sensitive to the degree of the posterior more advectages. The overall Board Ford at (2014), the ford the sensitive to the degree of the proposed based to compare locations and that during approximately one-to-ormal.

#### The effects of non-normality on statistical test

Parametric tests assuming normality (e.g., t-test & ANOVA) are often robust against non-normality; but depending on the type of non-normality (shape), parametric tests can have type I errors different (often greater) from alpha and also low power (increased type II error).

What happens if the Type I error probability (rate) is *greater* than alpha? i.e., increase number of False Positives.

22

#### The effects of non-normality on statistical test

Parametric tests assuming normality (e.g., t-test & ANOVA) are often robust against non-normality; but depending on the type of non-normality (shape), parametric tests can have type I errors different (often greater) from alpha and also low power (increased type II error).

What happens if the Type I error probability (rate) is *greater* than alpha? i.e., increase number of False Positives.

What happens if the Type I error probability (rate) is *smaller* than alpha? decrease False Positives but also decrease True Positives (i.e., lower statistical power).

23

## Type I versus Type II errors – the "common" view

A **Type I error (false positive)** is an **error** in every sense of the word. A conclusion is drawn that the null hypothesis is false when, in fact, it is true.

Therefore, **Type** I errors are generally considered more serious than **Type II** errors (false negatives).

Type II errors are often considered as "oh well, we were not able to detect an effect"...perhaps increase sample size!

Adapted from http://davidmlane.com/hyperstat/A2917.html

#### Type I versus Type II errors – the "common" view

A **Type I error (false positive)** is an **error** in every sense of the word. A conclusion is drawn that the null hypothesis is false when, in fact, it is true.

Therefore, **Type** I errors are generally considered more serious than **Type II** errors (false negatives). Type II errors are often considered as "oh well, we were not able to detect an effect"...perhaps increase sample size!

Adapted from http://davidmlane.com/hyperstat/A2917.html

When committing a type I error, you are stating that something that is false to be true.

CONFUSING: When committing at type II error, you are NOT stating that something that is true to be false (you are just not discovering something new).

25

Non-parametric tests based on ranks are those that can handle non-normal data

These are the main tests traditionally used in Biology for comparing samples:

1) For comparing two samples (analogue of the parametric two sample t-test) – *The Mann–Whitney U-test* (also known as the Mann–Whitney–Wilcoxon test, the Wilcoxon rank-sum test, or the Wilcoxon two-sample test).

26

## Non-parametric tests based on ranks are those that can handle non-normal data

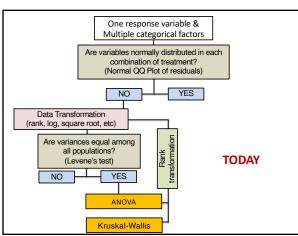
These are the main tests traditionally used in Biology for comparing samples:

1) For comparing two samples (analogue of the parametric two sample t-test) – *The Mann–Whitney U-test* (also known as the Mann–Whitney–Wilcoxon test, the Wilcoxon rank-sum test, or the Wilcoxon two-sample test).

2) For comparing multiple samples (analogue of the parametric ANOVA) – *The Kruskal-Wallis test* (generalization of the U-test)

The P-value for the *The Mann–Whitney U-test and the The Kruskal-Wallis test* is mathematically the same; as such, we will cover only the latter.

Note: remember that  $t^2 = F$ ; we often cover t-tests (and not only ANOVAs) in courses for two main reasons -[1] one sample t-tests; [2] understand the nature of post-hoc testing (e.g., post-hoc pairwise comparisons of means after ANOVA and because there is a t-test dealing with samples when their populations differ in their variances).



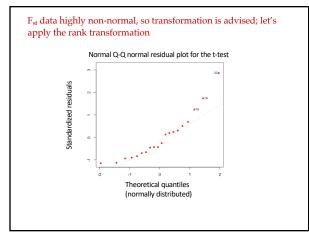


288 genetic polymorphism) in two populations of the American oyster, Crassestrat virginica. McDonald et al. (1996) collected data on Fg; for six anonymous DNA polymorphisms (variation in random bits of DNA on oh known function) and compared the Fg; values of the six DNA polymorphisms to Fg; values on 13 proteins from Buroker (1983). The biological question was whether protein polymorphisms would have generally lower or higher Fg; values than anonymous DNA polymorphisms. McDonald et al. (1996) knew that the theoretical distribution of Fg; for two populations is highly skewed, so they analyzed the data with a Kruskal-Wallis test.

|                         |                |            |             | ruskal–Wallis test starts by substituting  |
|-------------------------|----------------|------------|-------------|--|
| the rank in the overall | data set for e | ach measu  | rement v    | alue. The smallest value gets a rank of 1,   |
| the second-smallest ge  | ts a ran 🕅 a   | 1914. IBOI | nspar       | ametric tests are based on rank transformations  |
| two Fst values of -0.00 | are tied for   | second and | 1 third, so | they get a rank of 2.5.  |
|                         |                |            |             |  |
|                         |                |            | -           |  |
|                         | gene           | class      | FST         | Example: Fst is a measure of the amount of   |
|                         | CUIE           | DNIA       | 0.007       | the second dependence of the second dependence |

|  | Bene  | ciuss      | • 51                              | Example. 151 15 a filedulate of the amount of  |
|--|---|------------|-----------------------------------|--|
|  | CVJ5  | DNA        | -0.006                            | geographic variation in a genetic polymorphism.  |
|  | CVB1  | DNA        | -0.005                            | Here, McDonald et al. (1996) compared two  |
|  | 6Pgd  | protein    | -0.005                            | populations of the American oyster regarding the Fst   |
|  | Pgi   | protein    | -0.002                            |  |
|  | CVL3  | DNA        | 0.003                             | based on six anonymous DNA polymorphisms   |
|  | Est-3                                       | protein    | 0.004                             | (variation in random bits of DNA of no known   |
|  | Lap-2                                       | protein    | 0.006                             | function) and compared them to Fst values on 13  |
|  | Pgm-1                                       | protein    | 0.015                             | proteins.  |
|  | Aat-2                                       | protein    | 0.016                             | *  |
|  | Adk-1                                       | protein    | 0.016                             | Question: Do protein differ in FST values in contrast  |
|  | Sdh   | protein    | 0.024                             | to anonymous DNA polymorphisms?  |
|  | Acp-3                                       | protein    | 0.041                             | to anonymous DNA polymorphisms:  |
|  | Pgm-2                                       | protein    | 0.044                             | Zero Fst = no genetic variation (panmictic)  |
|  | Lap-1                                       | protein    | 0.049                             | negative Fst = more genetic variation within   |
|  | CVL1  | DNA        | 0.053                             | populations than between the two populations being   |
|  | Mpi-2                                       | protein    | 0.058                             | compared.  |
|  | Ap-1  | protein    | 0.066                             | •  |
|  | CVJ6  | DNA        | 0.095                             | positive Fst = more variation between populations than   |
|  | CVB2m                                       | DNA        | 0.116                             | within the two populations being compared.   |
|  | Est-1                                       | protein    | 0.163                             |  |
| her formidable formula                               | f the rank<br>that <b>itu</b> ps <b>i</b> ¢ | s for each | group, ti<br>sthaisc <b>ibo</b> e | en the test statistic, H. H is given by a<br>kacian/kesskithelitahlashumong Patps/rom McDonald et al. (1996) |
| vith an adjustment for the<br>hat the probability of |   |            |                                   |  |
| he P value correspond                                |   |            |                                   |  |
| roups minus 1. For th                                | ,   |            |                                   |  |
| otein is 10.68, H=0.0<br>pothesis that the Fs        |   |            |                                   |  |
| ected.   |   |            |                                   |  |
| For the reasons given                                |   |            |                                   |  |

For the reasons given anova. It gives a P value





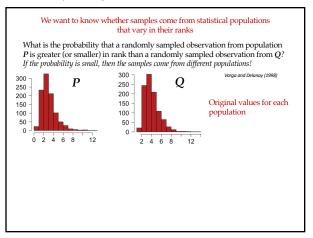
genetic polymorphism) in two populations of the American oyster, Crassestrat virginicat. McDonald et al. (1996) collected data on FST for six anonymous DNA polymorphisms (variation in random bits of DNA of no known function) and compared the FST values of the six DNA polymorphisms to FST values on 13 proteins from Buroker (1983). The biological question was whether protein polymorphisms. McDonald et al. (1996) know that the theorem anonymous DNA polymorphisms. McDonald et al. (1996) know that the theorem cital distribution of FST for two populations is highly skewed, so they analyzed the data with a Kruskal-Wallis test. When working will be americanous variable, the Kruskal-Wallis test. When working will be americanous variable, the Kruskal-Wallis test value gets a rank of 1, the second-smallest gets a rank of 1, of the second-smallest gets arout of the reaction of the second-smallest gets arout of the reaction of the reaction of the second-smallest gets arout of the reaction of the reaction of the second-smallest gets arout of the reaction of the second-smallest gets arout of the reaction of the reaction of the second-smallest gets arout of the reaction of the second-smallest gets arout of the reaction of the reaction of the second-smallest gets arout of the reaction of t

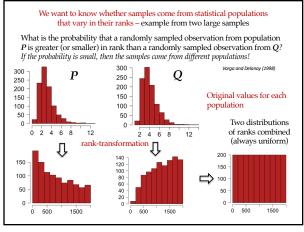
|       |         |        |      |      | sales based on transformations |
|-------|---------|--------|------|------|--------------------------------|
| gene  | class   | FST    | Rank | Rank |                                |
| CVJ5  | DNA     | -0.006 | 1    |      |                                |
| CVB1  | DNA     | -0.005 | 2.5  |      | (2+3)/2=2.5                    |
| 6Pgd  | protein | -0.005 |      | 2.5  | (2+3)/2-2.3                    |
| Pgi   | protein | -0.002 | -    | 4    |                                |
| CVL3  | DNA     | 0.003  | 5    |      |                                |
| Est-3 | protein | 0.004  |      | 6    |                                |
| Lap-2 | protein | 0.006  |      | 7    |                                |
| Pgm-1 | protein | 0.015  |      | 8    |                                |
| Aat-2 | protein | 0.016  |      | 9.5  | (0:10)/2.05                    |
| Adk-1 | protein | 0.016  |      | 9.5  | (9+10)/2=9.5                   |
| Sdh   | protein | 0.024  |      | 11   |                                |
| Acp-3 | protein | 0.041  |      | 12   |                                |
| Pgm-2 | protein | 0.044  |      | 13   |                                |
| Lap-1 | protein | 0.049  |      | 14   |                                |
| CVL1  | DNA     | 0.053  | 15   |      |                                |
| Mpi-2 | protein | 0.058  |      | 16   |                                |
| Ap-1  | protein | 0.066  |      | 17   |                                |
| CVJ6  | DNA     | 0.095  | 18   |      |                                |
| CVB2m | DNA     | 0.116  | 19   |      |                                |
| Est-1 | protein | 0.163  |      | 20   |                                |

rather formidable fo with an adjuster

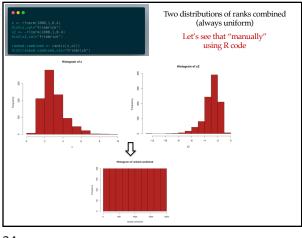
rather formidable to with an adjustment: 31that the probability 31the *P* value correspo groups minus 1. For protein is 10.68, H=( hypothesis that the 1 rejected.

For the reasons given above, I think it would actually be better to analyze the oyster data with one-way anova. It gives a P value of 0.75, which fortunately would not change the conclusions of McDonald et al.

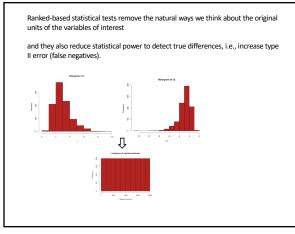




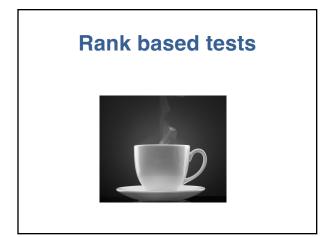


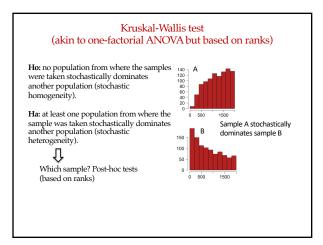




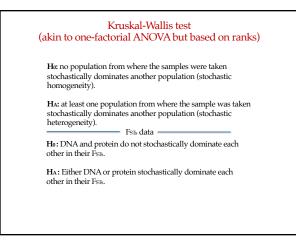


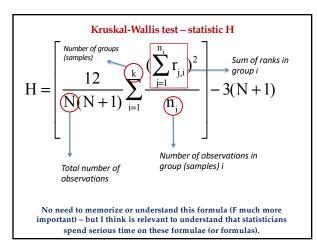




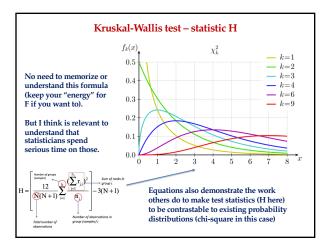






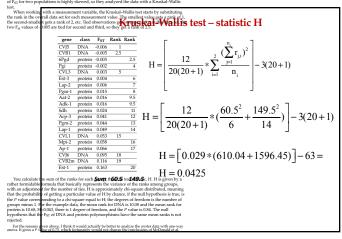


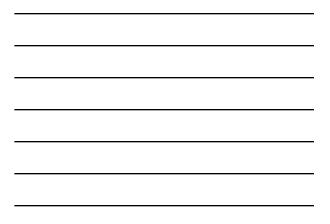






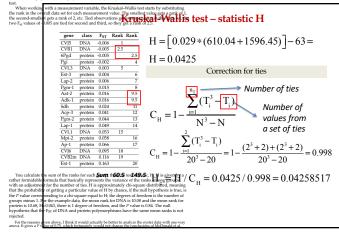
genetic polymorphism) in two populations of the American systex Cassastra inspirate. McDonald et al. (1986) collected data on F<sub>37</sub> for six anonymous DNA polymorphisms (variation in random bits of DNA of no harown function) and compared the F<sub>37</sub> values of the six DNA polymorphisms to F<sub>37</sub> values on 11 proteins from Barnier (1983). The biological question was whether protein polymorphisms world harge generally over or higher F<sub>3</sub> values fam anonymous DNA polymorphisms. McDanald et al. (1996) knew that the flowerical distribution of F<sub>37</sub> for two populations is highly slower, to show any advectional data with a Kraskal-Valin



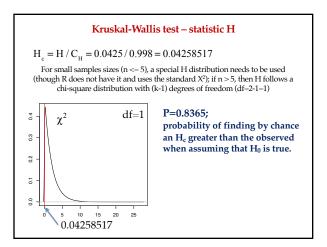


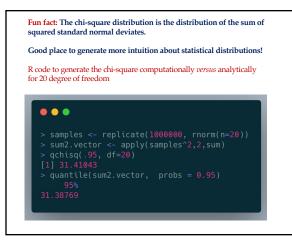


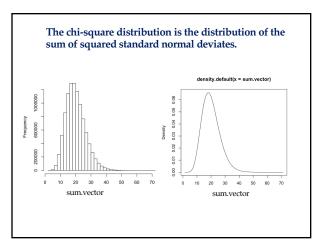
genetic polymorphism) in two populations of the American oyster, Crassestra ringinica, McDonald et al. (1996) collected data on F5 rfs vis. anonymous DNA polymorphisms (vriation in random bits of DNA of no kovory functiona) and compared the F57 values of the vis. DNA polymorphisms to F54 values and 15 proteins from Brancker (1983). The biological question was whether protein polymorphisms. McDonald et al. (1996) knew that the Branckeranonymous DNA polymorphisms. McDonald et al. (1996) knew that the Brancker-Holl of F54 or two polymorphisms is highly sheened to why a Aryadia data Walla



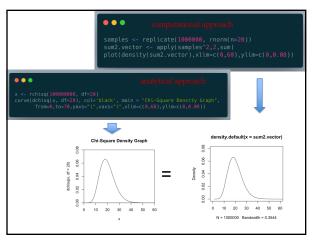








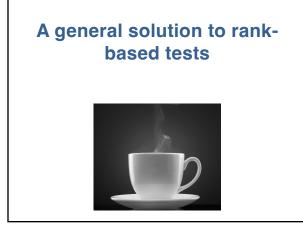






The chi-square distribution is the distribution of the sum of squared standard normal deviates.

*fun fact:* The F distribution is the ratio of two (scaled) chi-square distributed values. The scaling is done by appropriate division of degrees of freedom.



## Kruskal-Wallis test is equivalent (close enough) to an ANOVA on ranks Ho: no sample stochastically dominates another sample (stochastic homogeneity). Ha: at least one sample stochastically dominates one other sample (stochastic heterogeneity).

"Stochastic homogeneity is equivalent to the equality of the expected values of the rank sample means. This finding implies that the null hypothesis of stochastic homogeneity can be tested by an ANOVA performed on the rank transforms, which is essentially equivalent to doing a Kruskal-Wallis H test." Varga and Delanay (1998)

urnal of Educational and Behavioral Statistic mmer 1998, Vol. 23, No. 2, pp. 170–192

The Kruskal-Wallis Test and Stochastic Homogeneity András Vargha Eötvős Loránd University Harold D. Delaney University of New Mexico

49

### Kruskal-Wallis test = ANOVA on ranks

#### Kruskal-Wallis:

Ho: no sample stochastically dominates another sample (stochastic ĥomogeneity).

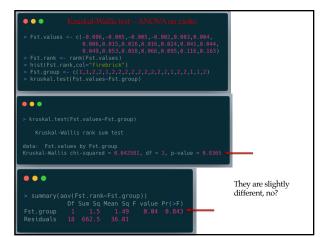
Ha: at least one sample stochastically dominates one other sample (stochastic heterogeneity).



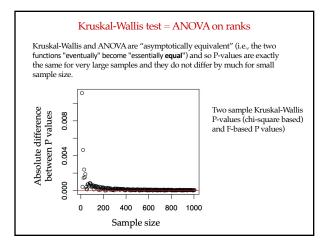
#### ANOVA:

Ho: no mean differences in ranked values

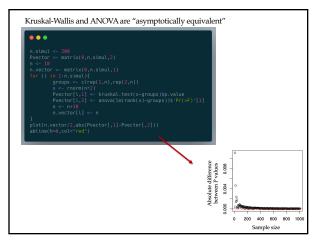
Ha: at least one sample differs in mean ranked values from another sample







Kruskal-Wallis and ANOVA are "asymptotically equivalent" and so P-values are the same for very large samples and they do not differ by much for small sample size. Using R code to demonstrate the asymptotic equivalence. n.simul <- 200 Pvector <- matrix(0,n.simul,2) n <- 10 n.vector <- matrix(0,n.simul,1) for (i in 1:n.simul){ groups <- c(rep(1,n),rep(2,n)) x <- rnorm(n\*2) Pvector[i,1] <- kruskal.test(x-groups)\$p.value Pvector[i,2] <- norao(lm(rank(x)-groups))\$'Pr(>F)'[1] n <- n + 10 n.vector[i] <- n } plot(n.vector/2,abs(Pvector[,1]-Pvector[,2])) abline(h=0,col="red")





#### Kruskal-Wallis test = ANOVA on ranks

Kruskal-Wallis and ANOVA are "asymptotically equivalent" and so P-values are exactly the same for very large samples and they do not differ by much for small sample size.

Because of the equivalence, we can then expand nonparametric analysis based on ranks to any multi-factorial ANOVAs, regressions, MANOVA, ANCOVA, etc

## 55

## **NOTE:** Non-parametric tests are those that can handle non-normal data

There is a common misunderstanding in the statistical literature and among practitioners, including many biostatistics books, that non-parametric tests can also handle differences in variances among samples.

THIS IS NOT TRUE! They are also affected by variance differences among groups/treatments (i.e., homoscedasticity).

Test variance differences in ranks (almost never done in the literature)!

